

Statistical Data Reduction via Construction of Sample Space Partitions*

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The statistical data reduction problem is presented as a problem of construction of sample space partitions. Then, the algorithm for synthesis of an ϵ -sufficient partition of a sample space is derived and its modification from the view-point of applications is formulated and discussed.

1. INTRODUCTION

Let the triple $(\Omega, \mathfrak{A}, P_\zeta)$ be a probability space: here Ω is a set whose elements are called ω 's, \mathfrak{A} denotes the σ -algebra of all subsets of Ω , P_ζ is a probability measure defined on the (measurable) space (Ω, \mathfrak{A}) . Let $\zeta(\omega)$ be the random variable corresponding to P_ζ and with range Ω . We shall call $(\Omega, \mathfrak{A}, P_\zeta)$ the *parameter space*. This name will be used also in referring simply to Ω .

We shall call $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ the *sample space*, where (X, \mathfrak{X}) is the measurable space of outcomes of an experiment and $P_{\xi|\omega}$ are conditional probability measures defined on (X, \mathfrak{X}) for each given parameter value $\omega \in \Omega$. Elements of the real space X are called x 's, \mathfrak{X} denotes the σ -algebra of all subsets of X . The set Ω can be also considered as an *index set* of probability measures $P_{\xi|\omega}$ on (X, \mathfrak{X}) . $\xi(\omega)$ is a random variable defined on the space Ω and taking its values in X . The name "sample space" is also used when referring only to X , its first element.

Let Y be a proper subset of X and let (Y, \mathfrak{Y}) be the measurable space with \mathfrak{Y} being the σ -algebra of all subsets of Y . We define the *problem of data reduction* as the problem of finding a partition \mathcal{A}_T of X defined by some measurable transformation T from (X, \mathfrak{X}) onto (Y, \mathfrak{Y}) . In other words, the problem of data reduction may be considered as the problem of searching the new experiment to be performed which is nothing different than the determination of a new random variable $\eta(\omega)$ defined

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on Ω which may be expressed as a following composition:

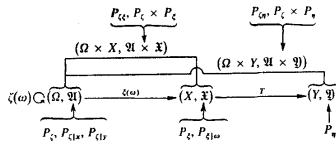
$$(1) \quad \eta(\omega) = T \circ \xi(\omega)$$

To each point $y \in Y$ corresponds some event $A_y \in \mathfrak{X}$ such that

$$(2) \quad Tx = y \quad \text{for all } x \in A_y$$

and, of course, by definition $A_y \in \mathcal{A}_T$.

The diagram below summarizes the principle notations to be used and gives the view of their relationships, where all the probability measures are generated in a standard way, provided $P_{\xi}, \{P_{\xi|\omega}, \omega \in \Omega\}$, T are given and η is defined by (1).



It should be clear that usually some constraints are imposed on a class of transformations to which T belongs. These are constraints concerning preservation under transformation T of information about the unknown value of parameter ω which is incorporated in events $A \in \mathfrak{X}$. To be able to make it more clear we introduce now some additional notations. Let D be an arbitrary space of actions or decisions d , let L be a loss function defined on $\Omega \times D$, let \mathcal{B} be a class of \mathfrak{X} -measurable decision functions δ with the range D . Further, let $\mathfrak{X}' \subset \mathfrak{X}$ be the σ -algebra generated by the partition \mathcal{A}_T and let $\mathcal{B}' \subset \mathcal{B}$ be a class of \mathfrak{X}' -measurable decision functions δ' . We are now in position to give the following definition.

Definition 1. The space X and the partition \mathcal{A}_T are said to be equally informative if there exists an element $\delta'_0 \in \mathcal{B}'$ such that

$$(3) \quad r(P_{\xi}, \delta'_0) = \inf_{\delta \in \mathcal{B}'} r(P_{\xi}, \delta)$$

where

$$(4) \quad r(P_{\xi}, \delta) = \int_{\Omega \times X} L(\omega, \delta(x)) dP_{\xi}$$

In the sequel we shall consider partitions \mathcal{A}_T which are "as informative as" X , as well as, such which are not. We remark that in general case only the latter lead to the essential data reduction. This statement is clarified later.

In Backwell and Girshick [1] may be found the following definition of a *sufficient partition*.

Definition 2. Let $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ be a sample space. A partition \mathcal{A} of X is said to be sufficient on $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ if for every bounded function f defined on X and every $A \in \mathcal{A}$, the conditional expectation of f , given A and ω

$$E_{\omega}(f | A) = \frac{1}{P_{\xi|\omega}(A)} \int_A f dP_{\xi|\omega}$$

is independent of ω for those $\omega \in \Omega$ for which $P_{\xi|\omega}(A) > 0$.

Using the factorization theorem (see [1] for the formulation and proof) one can prove

Theorem 1. Let \mathcal{A} be a sufficient partition on $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$. Then X and \mathcal{A} are equally informative.

It follows from Theorem 1 that if \mathcal{A} is a non-trivial sufficient partition (i.e. such a sufficient partition which does not exclusively consists of individual points of X), then instead of making precise measurements of the physical parameters of some objects represented by the vector $x \in X$, one can check only to which $A \in \mathcal{A}$ this vector belongs. If there exist non-trivial sufficient partitions \mathcal{A} of X , the question arises how to construct the minimal sufficient partition.

The appropriate algorithm may be readily written on the basis of Lemma 8.4.1 and Lemma 8.4.3 given in [1] under the following assumptions:

- (a) the sample space $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ is such that for each $x \in X$ there exists at least one $\omega \in \Omega$ with $P_{\xi|\omega}(x) > 0$,
- (b) the parameter space Ω is finite.

The assumption (a) means that the space X is such that its points really occur as results of the experiment performed. It is clear that from the view-point of applications the assumption (b) can not be considered as a restriction.

3. SUFFICIENT STATISTICS

Classically the sufficient statistic T on $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ is defined as a random variable such that the partition \mathcal{A}_T of X determined by T is sufficient on $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$.

Proposition 1. Let $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ be a sample space, let $(\Omega, \mathfrak{A}, P_{\xi})$ be a parameter space, and let T be a random variable defined on X and with range $Y \ni y$. Then T is a sufficient statistic on $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ if and only if for each pair (x, y) such

that

$$(5) \quad y = Tx$$

the equality

$$(6) \quad P_{\xi|x}(B) = P_{\xi|y}(B)$$

holds for all $B \in \mathfrak{A}$ such that $\int_B P_{\xi|\omega}(x) P_{\xi}(\omega) d\omega > 0$.

Proof. Suppose that T is a sufficient statistic on $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$ and \mathfrak{A}_T is the corresponding sufficient partition. Then for every $x \in A_y$ and each $A_y \in \mathfrak{A}_T$

$$Tx = y$$

Define

$$s_{\omega}(x) = \frac{P_{\xi|\omega}(x)}{\int_{\Omega} P_{\xi|\omega}(x) P_{\xi}(\omega) d\omega}$$

assuming that for each x the denominator is positive which is true by hypothesis (see Definition 2). This is also equivalent to the appropriate condition in Proposition 1. Then, from factorization theorem for sufficient statistics

$$s_{\omega}(x) = \frac{h(Tx, \omega) q(x)}{\int_{\Omega} h(Tx, \omega) q(x) P_{\xi}(\omega) d\omega} = \frac{h(Tx, \omega)}{\int_{\Omega} h(Tx, \omega) P_{\xi}(\omega) d\omega} = r_{\omega}(Tx)$$

and

$$(7) \quad P_{\xi|\omega}(x) = r_{\omega}(Tx) P_{\xi}(x).$$

Using (7) we obtain

$$(8) \quad \begin{aligned} P_{\xi|A}(B) &= \frac{1}{P_{\xi}(A)} \int_A dx \int_B P_{\xi|\omega}(x) P_{\xi}(\omega) d\omega = \\ &= \frac{1}{P_{\xi}(A)} \int_A P_{\xi}(x) dx \int_B r_{\omega}(Tx) P_{\xi}(\omega) d\omega \end{aligned}$$

where $A \in \mathfrak{X}$, $B \in \mathfrak{A}$. $P_{\xi|A}(B)$ may be also expressed as

$$(9) \quad P_{\xi|A}(B) = \frac{1}{P_{\xi}(A)} \int_A P_{\xi|x}(B) dP_{\xi}(x) = \frac{1}{P_{\xi}(A)} \int_A P_{\xi}(x) dx \int_B P_{\xi|x}(\omega) d\omega$$

Now, assuming that $A = A_y \in \mathfrak{A}_T$ we conclude that for $x \in A_y$ $r_{\omega}(Tx)$ does not depend upon x . Moreover, the left hand sides of (8) and (9) become $P_{\xi|y}(B)$. Next, comparing the right hand sides of (8) and (9) we conclude that $\int_B P_{\xi|x}(\omega) d\omega$ does not

446 depend upon $x \in A_y$. This means that (6) holds for all $x \in A_y$ and $B \in \mathfrak{A}$ such that $\int_B P_{\xi|\omega}(x) P_{\xi}(\omega) d\omega > 0$. Conversely, suppose (6) together with (5) holds. Since

$$P_{\xi|x}(\omega) = \frac{P_{\xi|\omega}(x) P_{\xi}(\omega)}{P_{\xi}(x)}$$

and

$$P_{\xi|y}(\omega) = \frac{P_{\eta|\omega}(y) P_{\xi}(\omega)}{P_{\eta}(y)}$$

we obtain

$$P_{\xi|\omega}(x) = \frac{P_{\eta|\omega}(y)}{P_{\eta}(y)} P_{\xi}(x)$$

or

$$P_{\xi|\omega}(x) = \frac{P_{\eta|\omega}(Tx)}{P_{\eta}(Tx)} P_{\xi}(x)$$

which making appropriate definitions is equivalent to the necessary and sufficient condition (given by factorization theorem) for a random variable T to be a sufficient statistic.

This completes the proof of the proposition.

4. PARTITIONS WHICH ARE ε -SUFFICIENT ON A SAMPLE SPACE. GENERAL CONSIDERATIONS

Let T be any measurable transformation from the measurable space (X, \mathfrak{X}) onto a measurable space (Y, \mathfrak{Y}) as stated in Introduction. We have by definition

$$(10) \quad P_{\xi|y}(B) = \frac{P_{\xi\eta}(B, y)}{P_{\eta}(y)} = \frac{P_{\xi\zeta}(B, A)}{P_{\zeta}(A)}, \quad B \in \mathfrak{Y}, \quad A \in \mathfrak{X}$$

where

$$(11) \quad A = T^{-1}y = \{x : Tx = y\}.$$

From (10)

$$(12) \quad P_{\xi|y}(B) = \frac{P_{\xi\zeta}(B, A)}{P_{\eta}(y)} = \frac{1}{P_{\eta}(y)} \int_A P_{\xi|x}(B) dP_{\xi}$$

which is equivalent to (6), provided A is an element of a sufficient partition. Note that $[P_{\eta}]$ the following relations hold:

$$(13) \quad \frac{P_{\xi}(x)}{P_{\eta}(y)} \geq 0, \quad x \in X, \quad y \in Y$$

$$(14) \quad \frac{1}{P_{\eta}(y)} \int_A dP_{\xi} = 1, \quad A \in \mathfrak{X} \quad \text{and} \quad A = T^{-1}y.$$

Let h be any concave function. Then, since (13) and (14) hold, we can apply to (12) Jensen's inequality. We get then

$$(15) \quad h\left(\frac{1}{P_{\eta}(y)} \int_A P_{\xi|x}(B) dP_{\xi}\right) \geq \frac{1}{P_{\eta}(y)} \int_A h(P_{\xi|x}(B)) dP_{\xi}.$$

Taking the integral of the both sides of (15) over the space $\Omega \times Y$, we obtain

$$(16) \quad \int_{\Omega} d\omega \int_Y P_{\eta}(y) h(P_{\xi|y}(\omega)) dy \geq \int_{\Omega} d\omega \int_X P_{\xi}(x) h(P_{\xi|x}(\omega)) dx.$$

If we denote

$$\varepsilon' = \int_{\Omega} d\omega \int_Y P_{\eta}(y) h(P_{\xi|y}(\omega)) dy - \int_{\Omega} d\omega \int_X P_{\xi}(x) h(P_{\xi|x}(\omega)) dx$$

and if we choose

$$(17) \quad h(P) = -P \log P$$

where the basis of logarithms is equal to 2, then, from (16) we get

$$(18) \quad - \int_{\Omega} d\omega \int_Y P_{\eta}(\omega, y) \log P_{\xi|y}(\omega) dy + \\ + \int_{\Omega} d\omega \int_X P_{\xi}(\omega, x) \log P_{\xi|x}(\omega) dx \leq \varepsilon$$

where $\varepsilon \geq \varepsilon'$ is a non-negative number.

Definition 3. The measurable transformation T from the measurable space (X, \mathfrak{X}) onto a measurable space (Y, \mathfrak{Y}) is ε -sufficient if the inequality (18) holds with $\varepsilon > 0$. The partition \mathcal{A}_T of X determined by such T is said to be ε -sufficient on $(X, \mathfrak{X}, \Omega, P_{\xi|\omega})$.

Remark 1. It is obvious that if and only if $\varepsilon = 0$ in (18) then T is sufficient.

The concept of ε -sufficient transformation was in the different way first introduced by Perez [2] and was studied by him in [3, 4, 5]. The equivalence of both definitions is straightforward.

Equality (3) defines the Bayes risk. As stated by Theorem 1 the Bayes risk remains unchanged if X is replaced by its sufficient partition \mathcal{A}_T . This is, however, no longer the case if \mathcal{A}_T is an ε -sufficient partition. This means that in the case of ε -sufficient data reduction we do need an estimate of the Bayes risk increase. Such an estimate

448 is given by the Perez's theorem to be found in [3]. See also Perez [4, 5] for further considerations.

5. ε -SUFFICIENT DATA REDUCTION. CONSTRUCTIVE RESULTS

The class of sample spaces to be considered in this section is that with finite parameter space Ω ; a member of this class is denoted by $(X, \mathfrak{X}, \Omega, P_{\xi|\omega_i})$. Let M be the number of points ω_i in Ω . We will give an algorithm for constructing an ε -sufficient partition. This partition turns out to be finite. The considerations of this section extend the earlier results of the author presented in [6] and [7].

Let us assume, for a moment without any motivation, that in the case of finite parameter space it is possible for any positive value ε to construct an ε -sufficient partition of X which is finite. This implies the finiteness of the space Y . Let K be the number of elements A_j in \mathcal{A}_T (or the number of elements y_j in Y). With these assumptions we can replace the inequality (18) by

$$(19) \quad - \sum_{\Omega} \sum_{j=1}^K P_{\xi|\omega}(\omega, A_j) \log P_{\xi|A_j}(\omega) + \sum_{\Omega} \int_X P_{\xi|x}(\omega) \log P_{\xi|x}(\omega) dP_{\xi}(x) \leq \varepsilon.$$

Now, we give an information-theoretic interpretation of the problem of searching an ε -sufficient transformation in the case considered. Let $(\Omega, X, \mathfrak{X}, P_{\xi|\omega_i})$ and $(\Omega, \mathfrak{A}, P_{\xi})$ be a semicontinuous channel and a source of information, respectively. The average amount of information per transmission received through the channel is given by

$$R = H(\Omega) - H(\Omega | X)$$

where

$$H(\Omega) = - \sum_{\Omega} P_{\xi}(\omega) \log P_{\xi}(\omega),$$

$$H(\Omega | X) = - \sum_{\Omega} \int_X P_{\xi|x}(\omega) \log P_{\xi|x}(\omega) dP_{\xi}(x).$$

It was proved by Feinstein [8] that for any $\varepsilon > 0$ one can replace the semicontinuous channel defined above by a discrete one which assures the decrease of the average amount of information per transmission not greater than ε . This means that one can find the transformation T defined above for which

$$(20) \quad R - R' \leq \varepsilon$$

with

$$R' = H(\Omega) - H(\Omega | \mathcal{A}_T)$$

where

$$H(\Omega | \mathcal{A}_T) = -\sum_{\Omega} \sum_{j=1}^K P_{\zeta_i}(\omega, A_j) \log P_{\zeta_i A_j}(\omega).$$

Clearly, (20) is equivalent to (19). This leads us to the following assertion:

Theorem 2. *If the parameter space Ω is finite, then for any positive ε there exists an ε -sufficient measurable transformation T from the measurable space (X, \mathfrak{X}) onto a finite measurable space (Y, \mathfrak{Y}) . This means that the corresponding ε -sufficient partition of X is finite.*

We propose an algorithm for constructing an ε -sufficient partition \mathcal{A}_T of X . This algorithm is based on the proof of Feinstein's theorem, given in [8].

To construct an appropriate set $\mathcal{A}_T = \{A_j\}$ we explicitly put

$$(21) \quad -\log P_{\zeta_i X}(\omega_i) = 0 \quad \text{for} \quad P_{\zeta_i X}(\omega_i) = 0, \quad \omega_i \in \Omega.$$

Then, denoting by m any positive integer, we define the following sets:

$$(22) \quad A_m(\omega_i) = \{x : -\log P_{\zeta_i X}(\omega_i) < m\}, \quad \omega_i \in \Omega,$$

$$(23) \quad A_m = \bigcap_{\Omega} A_m(\omega_i).$$

Algorithm 1

1° Find the smallest subscript $m = m_0$ such that

$$(24) \quad -\sum_{\Omega} P_{\zeta_i}(\omega_i, X \setminus A_{m_0}) \log P_{\zeta_i}(\omega_i, X \setminus A_{m_0}) \\ + P_{\zeta_i}(X \setminus A_{m_0}) \log P_{\zeta_i}(X \setminus A_{m_0}) \leq \gamma$$

where

$$(25) \quad P_{\zeta_i}(\omega_i, X \setminus A_{m_0}) = P_{\zeta_i|\omega_i}(X \setminus A_{m_0}) P_{\zeta_i}(\omega_i),$$

$$(26) \quad P_{\zeta_i}(X \setminus A_{m_0}) = \sum_{\Omega} P_{\zeta_i}(\omega_i, X \setminus A_{m_0})$$

and γ is a positive number such that

$$(27) \quad \gamma < \varepsilon$$

where ε is a positive number chosen before.

As a result of this step of the algorithm we obtain the set A_{m_0} .

2° Choose the smallest positive integer n such that

$$(28) \quad \frac{1}{n} P_{\zeta_i}(A_{m_0}) \leq \varepsilon - \gamma.$$

450 3° Find

$$(29) \quad P_{\max} = \sup_{\Omega \times A_{m_0}} P_{\zeta|x}(\omega_i)$$

and then

$$(30) \quad k_{\min} = [1 - n \log P_{\max}].$$

4° For each $\omega_i \in \Omega$ construct the following sequence of sets

$$(31) \quad A_{k,\omega_i} = \{x : 2^{-k/n} < P_{\zeta|x}(\omega_i) \leq 2^{-(k-1)/n}\} \cap A_{m_0}$$

$$k = k_{\min}, \dots, nm_0,$$

$$(32) \quad A_{0,\omega_i} = \{x : P_{\zeta|x}(\omega_i) = 0\} \cap A_{m_0}.$$

5° Construct all the following sets:

$$(33) \quad A(k_1, k_2, \dots, k_M) = \bigcap_{i=1}^M A_{k_i, \omega_i}$$

where

$$(34) \quad k_i = 0, k_{\min}, k_{\min} + 1, \dots, nm_0.$$

It is proved in the sequel that as a result of this step we obtain the set

$$(35) \quad \mathcal{A}_T = \{A_j\} = \{A(k_1, k_2, \dots, k_M)\}, \quad X \setminus A_{m_0}.$$

Now we make some remarks which will be found helpful in proving that the formulated algorithm possesses the desired properties.

Remark 2. The sequence A_m is a non-decreasing sequence of sets. Therefore

$$(36) \quad \lim A_m = \bigcup_{m=1}^{\infty} A_m = X.$$

Remark 3. It follows from relations (29), (30), (31) and (32) that

$$(37) \quad \bigcup_{k=0, k_{\min}, \dots, nm_0} A_{k,\omega_i} = A_{m_0}, \quad \omega_i \in \Omega,$$

$$(38) \quad \bigcup_{k=0, k_{\min}, \dots, nm_0} A_{k,\omega_i} = \emptyset, \quad \omega_i \in \Omega.$$

The same assertions are clearly true for the sets $A(k_1, k_2, \dots, k_M)$.

Remark 4. The set \mathcal{A}_T given by (35) is a partition of the space X (it is motivated by Remarks 2 and 3).

Remark 5. Since $0 < \gamma < \varepsilon$ and γ may assume any value from this interval, then for the fixed value ε we can obtain uncountably many partitions \mathcal{A}_T of the space X .

We formulate now the theorem concerning Algorithm 1.

Theorem 3. Let $(X, \mathfrak{X}, \Omega, P_{\xi|\omega_i})$ be a given sample space with the parameter space Ω which consists of M points, and let also the a priori probability measure $P_{\xi}(\omega)$ be given. Then a partition $\mathcal{A}_T = \{A_j\}$ of the space X obtained as a result of Algorithm 1 is ε -sufficient.

Proof. Taking into account Remark 2 and $\lim_{a \rightarrow 0} a \log a = 0$ we conclude that it is possible to find $m = m_0$ such that for the chosen positive $\gamma < \varepsilon$ and any $\varepsilon < 0$ we will have

$$(39) \quad \begin{aligned} & - \sum_{\omega_i \in \Omega} P_{\xi|\omega_i}(\omega_i, X \setminus A_{m_0}) \log P_{\xi|X \setminus A_{m_0}}(\omega_i) = \\ & = - \sum_{\omega_i \in \Omega} P_{\xi|\omega_i}(\omega_i, X \setminus A_{m_0}) \log P_{\xi|\omega_i}(\omega_i, X \setminus A_{m_0}) \\ & \quad + P_{\xi}(X \setminus A_{m_0}) \log P_{\xi}(X \setminus A_{m_0}) \leq \gamma \end{aligned}$$

where the set A_{m_0} is defined in the step 1° of Algorithm 1.

If $r > 0$, then

$$(40) \quad 2^{-r/n} < P_{\xi|x}(\omega_i) \leq 2^{-(r-1)/n}$$

for all x belonging to any $A(k_1, k_2, \dots, k_M)$ for which $k_i = r$, where n is defined by (28) and the sets $A(k_1, k_2, \dots, k_M)$ are defined by (33). Since

$$(41) \quad P_{\xi|A(k_1, k_2, \dots, k_M)}(\omega_i) = \frac{1}{P_{\xi}(A(k_1, k_2, \dots, k_M))} \int_{A(k_1, k_2, \dots, k_M)} P_{\xi|x}(\omega_i) dP_{\xi}$$

the same inequality should be true for $P_{\xi|A(k_1, k_2, \dots, k_M)}(\omega_i)$ except those cases when $P_{\xi}(A(k_1, k_2, \dots, k_M)) = 0$, which means that $P_{\xi|\omega_i}(\omega_i, A(k_1, k_2, \dots, k_M)) = 0$.

Further, if $k_i = 0$, then

$$(42) \quad P_{\xi|\omega_i}(\omega_i, A(k_1, k_2, \dots, k_M)) = 0$$

for all corresponding sets $A(k_1, k_2, \dots, k_M)$.

Define on A_{m_0} for each ω_i the following function (recall Remark 3):

$$(43) \quad g(\omega_i, x) = \begin{cases} -\log P_{\xi|A(k_1, k_2, \dots, k_M)}(\omega_i) & \text{if } x \in A(k_1, k_2, \dots, k_M) \\ & \text{such that } P_{\xi|\omega_i}(\omega_i, A(k_1, k_2, \dots, k_M)) > 0, \\ \text{any value at all other points of } A_{m_0}. \end{cases}$$

Thus

$$(44) \quad |-\log P_{\xi|x}(\omega_i) - g(\omega_i, x)| \leq \frac{1}{n} [P_{\xi|\omega_i}]$$

452 on A_{m_0} , $i = 1, 2, \dots, M$. Since $-\log P_{\zeta|x}(\omega_i)$ and $g(\omega_i, x)$ are positive $[P_{\zeta\zeta}]$, we have

$$(45) \quad \int_{A_{m_0}} \log P_{\zeta|x}(\omega_i) dP_{\zeta\zeta}(\omega_i, x) \geq - \int_{A_{m_0}} g(\omega_i, x) dP_{\zeta\zeta}(\omega_i, x) - \frac{1}{n} P_{\zeta\zeta}(\omega_i, A_{m_0}).$$

But

$$(46) \quad \int_{A_{m_0}} \varrho(\omega_i, x) dP_{\zeta\zeta}(\omega_i, x) = - \sum_{(k_1, k_2, \dots, k_M)} P_{\zeta\zeta}(\omega_i, A(k_1, k_2, \dots, k_M)) \log P_{\zeta|A(k_1, k_2, \dots, k_M)}(\omega_i).$$

So that, taking into account the definition of conditional entropy $H(\Omega | X)$ given above, and Eqs. (45) and (46), we obtain

$$(47) \quad H(\Omega | X) \geq - \sum_{i=1}^M \int_{A_{m_0}} \log P_{\zeta|x}(\omega_i) dP_{\zeta\zeta}(\omega_i, x) \geq - \sum_{i=1}^M \left\{ \sum_{(k_1, \dots, k_M)} P_{\zeta\zeta}(\omega_i, A(k_1, k_2, \dots, k_M)) \log P_{\zeta|A(k_1, k_2, \dots, k_M)}(\omega_i) - \frac{1}{n} P_{\zeta\zeta}(\omega_i, A_{m_0}) \right\}.$$

Now, taking into account (28) and (39), we obtain from (47) the following inequalities:

$$(48) \quad H(\Omega | X) \geq - \sum_{i=1}^M \sum_{(k_1, \dots, k_M)} P_{\zeta\zeta}(\omega_i, A(k_1, k_2, \dots, k_M)) \log P_{\zeta|A(k_1, k_2, \dots, k_M)}(\omega_i) - \frac{1}{n} P_{\zeta}(A_{m_0}) - \sum_{i=1}^M P_{\zeta\zeta}(\omega_i, X \setminus A_{m_0}) \log P_{\zeta|X \setminus A_{m_0}}(\omega_i) - \gamma \geq - \sum_{i=1}^M \sum_{A_j \in \mathcal{A}_T} P_{\zeta\zeta}(\omega_i, A_j) \log P_{\zeta|A_j}(\omega_i) - \varepsilon$$

where \mathcal{A}_T is defined by (35). One can very easily see that the inequality (48) is equivalent to the inequality (20). This proved ε -sufficiency of the partition \mathcal{A}_T of X .

This completes the proof of the theorem.

One can easily see that since $0 < \gamma < \varepsilon$ there exists some optimal γ which minimizes

$$(49) \quad \Delta = nm_0 - [1 - n \log P_{\max}],$$

i.e., γ corresponding to the minimal amount of the computational work to be done in steps 4° and 5° of Algorithm 1. The problem of this optimization, however, is in fact not very important from the view-point of applications of ε -sufficient data reduction.

In practical cases one will:

1. Restrict the computation to some bounded space X .
2. Assume some "regular" partition $\mathcal{S} = \{S\}$ of the space X , where the sets $S \in \mathcal{S}$ are "sufficiently small".
3. Assign the values $P_{\zeta|S}(\omega_i)$, $i = 1, 2, \dots, M$, found experimentally, to all points $x \in S$. This means that one will have in X only points x such that either $P_{\zeta|x}(\omega_i) = 0$ or $P_{\zeta|x}(\omega_i) > 0$ with the condition

$$(50) \quad P_{\zeta|x}(\omega_i) > \varrho$$

fulfilled for every x and each ω_i for which $P_{\zeta|x}(\omega_i) \neq 0$, where $\varrho > 0$ is small.

4. Construct $\mathcal{A}_T = \{A_j\}$, where every $A_j \in \mathcal{A}_T$ will consist of at least one set $S \in \mathcal{S}$, according to Algorithm 2, which is given below and is the obvious modification of Algorithm 1.

Algorithm 2

- 1° Find the smallest positive value of $P_{\zeta|S}(\omega_i)$, $i = 1, 2, \dots, M$,

$$P_{\min} = \inf_{\substack{S \in \mathcal{S} \\ P_{\zeta|S}(\omega_i) \neq 0}} P_{\zeta|S}(\omega_i)$$

and then choose

$$m_0 = \lceil -\log P_{\min} \rceil + 1.$$

- 2° Choose the smallest positive integer n such that

$$\frac{1}{n} \leq \varepsilon$$

where ε is a positive number chosen before.

- 3° Proceed as in Algorithm 1 with x replaced by S .

4° Proceed as in Algorithm 1 with x replaced by S and the operations of intersection with A_{m_0} deleted.

- 5° Proceed as in Algorithm 1.

As a result of Algorithm 2 one obtains a partition

$$\mathcal{A}_T = \{A(k_1, k_2, \dots, k_M)\}.$$

We remark that such a partition will be ε -sufficient with respect to the computed probability measures $P_{\zeta|S}(\omega_i)$, $i = 1, 2, \dots, M$, $S \in \mathcal{S}$.

Even if $\{P_{\zeta|\omega_i}\}$ and P_{ζ} are exactly known one can assume some "convenient" partition \mathcal{S} of X and compute "how sufficient" is this partition, i.e., one can compute the number

$$\begin{aligned} \varepsilon_{\mathcal{S}} = & - \sum_{i=1}^M \sum_{S_j \in \mathcal{S}} P_{\zeta\zeta^c}(\omega_i, S_j) \log P_{\zeta|\mathcal{S}}(\omega_i) \\ & + \sum_{i=1}^M \int_X \log P_{\zeta|x}(\omega_i) dP_{\zeta\zeta^c}(\omega_i, x) \end{aligned}$$

Then, if a partition \mathcal{A}_T should be ε -sufficient one can find

$$\varepsilon_{\mathcal{A}_T|\mathcal{S}} = \varepsilon - \varepsilon_{\mathcal{S}}$$

provided $\varepsilon > \varepsilon_{\mathcal{S}}$. Further, using Algorithm 2 one can find \mathcal{A}_T being an $\varepsilon_{\mathcal{A}_T|\mathcal{S}}$ -sufficient partition with respect to \mathcal{S} and being ε -sufficient on $(X, \mathfrak{X}, \Omega, P_{\zeta|\omega_i})$.

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Statistická redukce dat pomocí konstrukce rozkladů výběrového prostoru

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V článku je problém statistické redukce dat formulován jako problém konstrukce rozkladů výběrového prostoru. Uvažují se suficientní a ε -suficientní rozklady. Je uvedena nová definice suficientní statistiky, ze které plyne definice ε -suficientního rozkladu, jež je ekvivalentní definici Perezově. Nová definice suficientní statistiky umožnila dokázat, že pro konečný parametrový prostor je problém syntézy ε -suficientního rozkladu ekvivalentní problému redukce polospojitého kanálu na diskrétní kanál, nemá-li pokles průměrné informace na přenos překročit ε . To umožnilo odvodit algoritmus pro syntézu ε -suficientního rozkladu výběrového prostoru inspirovaný prací Feinsteina [8]. Je předložena a studována modifikace tohoto algoritmu.

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